Parametric Fitting

Kenichi Kanatani (interpreted by I. Esteban)

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Outline

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1 Introduction

2 General Theory

- Definitions
- Maximum Likelihood Estimation
- Covariance Matrix
- Hypothesis and Noise Level
- **3** Fitting for Image Points

4 Questions.... for you

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What's all about?

... fit geometric objects to multiple instances of another geometric object in an optimal manner ...

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Example



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Parametric Fitting



Fitting as Maximum Likelihood

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Fitting as Maximum Likelihood

Obtain covariance of estimation and residual

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- Fitting as Maximum Likelihood
- Obtain covariance of estimation and residual
- Hypothesis testing

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Fitting as Maximum Likelihood

- Obtain covariance of estimation and residual
- Hypothesis testing
- Examples

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Definitions Maximum Likelihood Estimation Covariance Matrix Hypothesis and Noise Level

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Object, data and relationship

- We want to fit an object given some data
- Let u be the vector that represents the object (n'-dim manifold - parameter space U)
- Let a₁...a_N be N vectors of the same object (m'-dim manifold - data space A)
- All a_{α} satisfy the same relation with u
- We want to find *u* optimally wrt the relation

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Noise

• a_{α} is observed in the presence of noise

$$a_{\alpha} = \bar{a}_{\alpha} + \Delta a_{\alpha}, \alpha = 1, ..., N.$$

- Δa_{α} is independent, zero mean and with covariance $\bar{V}[a_{\alpha}]$
- To a first approximation, Δa_{α} lives in the tangent space $T_{\bar{a}_{\alpha}}(A)$

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Constraints

- The true value \bar{a}_{α} and u are related by L constraints $F^{(k)}(\bar{a}_{\alpha}, u) = 0, k = 1, ..., L.$
- This is called the hypothesis and its assumed to be nonsingular (p.132)
- The rank of the hypothesis is the number of independent equations
- The *rank* is the codimension of *S* (manifold defined by the L eqs.)
- *S* is called the *geometric model* of the hypothesis

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Singularity

- Linear subspace of the TRUE value $\bar{V}_{\alpha} = \{P^{A}_{\bar{a}_{\alpha}} \nabla_{a} F^{(1)}(\bar{a}_{\alpha}, u), ..., P^{A}_{\bar{a}_{\alpha}} \nabla_{a} F^{(L)}(\bar{a}_{\alpha}, u)\}_{L} \in R^{m}$
- Linear subspace of the MEASURED value $V_{\alpha} = \{P_{a_{\alpha}}^{A} \nabla_{a} F^{(1)}(a_{\alpha}, u), ..., P_{a_{\alpha}}^{A} \nabla_{a} F^{(L)}(a_{\alpha}, u)\}_{L} \in R^{m}$
- In general, the rank of the hypothesis r coincides with the dimension l of the linear subspace \bar{V}_{α}
- If l < r then a_{α} is a singular datum
- If the dimension of V_{α} is larger than the dimesion of \bar{V}_{α} the hypothesis is degenerate

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Example??

- Image points
- Directions in space
- Motion estimation??

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Correction

- Assume a value for u (what value?)
- Correct (optimally) data to satisfy the hypothesis
 - Find Δa_{lpha} such that $ar{a}_{lpha} = a_{lpha} \Delta a_{lpha}$
 - This is the optimization:

$$J_{lpha} = (\Delta a_{lpha}, ar{V}[a_{lpha}]^{-} \Delta a_{lpha}) o {\it min}$$

Solution is given by:

$$\Delta a_{\alpha} = \bar{V}[a_{\alpha}] \sum_{k,l=1}^{L} \bar{W}_{\alpha}^{(kl)}(u) F^{(k)}(a_{\alpha}, u) \nabla a F^{(l)}(\bar{a}_{\alpha}, u)$$

• The residual \hat{J}_{α} is given by substituting one into another

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Estimation

- The probability density for all Δa_{α} is: $\left(\prod_{\beta=1}^{N} \frac{1}{\sqrt{(2\pi)^{r_{\beta}} |\bar{V}[a_{\beta}]|_{+}}}\right) e^{-\sum_{\alpha=1}^{N} (\Delta a_{\alpha}, \bar{V}[a_{\alpha}]^{-} \Delta a_{\alpha})/2}$
- This is the likelihood of the observed values Δa_{α} ??
- Given the residual \hat{J}_{α} this takes the form:

$$\left(\prod_{eta=1}^{\mathsf{N}}rac{1}{\sqrt{(2\pi)^{r_eta}|ar{m{V}}[m{a}_eta]|_+}}
ight)e^{-\sum_{lpha=1}^{\mathsf{N}}\hat{J}_lpha/2}$$

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Estimation 2

We now want to find u that maximizes that likelyhood (or minimizes the sum of residuals)

$$\bar{J}[u] = \sum_{\alpha=1}^{N} \hat{J}_{\alpha} \to min$$

This takes the full form:

$$\bar{J}[u] = \sum_{\alpha=1}^{N} \sum_{k,l=1}^{L} \bar{W}_{\alpha}^{(kl)}(u) F^{(k)}(a_{\alpha}, u) F^{(l)}(a_{\alpha}, u) \rightarrow min$$

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Practical Considerations

- $ar{W}^{(kl)}_{lpha}$ (pseudo inverse of V_{lpha}) depends on the true value $ar{a}_{lpha}$
- We approximate: $\bar{W}_{\alpha}^{(kl)} \approx \left((\nabla a F^{(k)}(a_{\alpha}, u), V[a_{\alpha}] \nabla F^{(l)}(a_{\alpha}, u) \right)_{r}^{-}$
- Thus:

$$J[u] = \sum_{\alpha=1}^{N} \sum_{k,l=1}^{L} W_{\alpha}^{(kl)}(u) F^{(k)}(a_{\alpha}, u) F^{(l)}(a_{\alpha}, u) \rightarrow min$$

• Yields the optimal estimate \hat{u} by numerical computation

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- We define the optimal estimation as maximum likelihood
- We obtain an optimal estimate by approximating the true covariance with the measured covariance
- The optimal estimate û is a random variable since it was obtained from noisy data
- We now study its behavior

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Some facts and definitions

- We define \bar{u} as the true value and u the random variable
- \bar{u} satisfies $F^{(k)}(\bar{a}_{\alpha}, u) = 0, k = 1, ..., L$
- Since its a random variable, its disturbed by noise: $u = \overline{u} + \Delta u$
- Δu is to a first approx contained in the tangent space

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Covariance of \hat{u}

We start by introducing the random variables a_α and u in the constraints: $F^{(k)}(a_α, u) = (∇aF^(k)_α, Δa_α) + (∇uF^(k)_α, Δu_α) + O(Δa_α, Δu)²$

Also:
$$\bar{W}^{(kl)}_{\alpha}(u) = \bar{W}^{(kl)}_{\alpha}(\bar{u}) + O(\Delta u)$$

After some magical math, we finally obtain:

$$\bar{V}[\hat{u}] = \left(\sum_{\alpha=1}^{N}\sum_{k,l=1}^{L}\bar{W}_{\alpha}^{(kl)}(\bar{u})(P_{\bar{u}}^{U}\nabla u\bar{F}_{\alpha}^{(k)})(P_{\bar{u}}^{U}\nabla u\bar{F}_{\alpha}^{(l)})^{T}\right)^{-1}$$

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Approximation

- Since the true value \bar{u} is used, we cannot compute it
- We approximate using the optimal estimate \hat{u} and the corrected data $\hat{a}_{\alpha} = a_{\alpha} \Delta a_{\alpha}$
- Alternatively, one approximation can be made using the measured data and not the corrected one.

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- The estimation is based on the *hypothesis* that the data {a_α} are random deviations from the true data {ā_α}
- The true data satisfies the constraints $F^{(k)}$
- Minimizing the sum of residuals means choosing û so that the hypothesis is most likely

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Hypothesis testing

- If the hypothesis is correct, the residual $\overline{J}[\hat{u}]$ should be zero for the true values
- This is generally NOT the case
- The bigger the residual, the lest likely the hypothesis is correct
- If the residual is much larger than expected according to the noise in the data {a_α} the hypothesis can be rejected
- To formalize this, we assume Gaussian noise in the data

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Strong hypothesis

- We want to reject a strong hypothesis
- A strong hypothesis is based on the residual of the TRUE value \bar{u}
- So, we consider the residual $\bar{J}[\bar{u}]$ and we let $\Delta u = 0$

This is given by:

$$\bar{J}[\bar{u}] = \sum_{\alpha=1}^{N} \sum_{k,l=1}^{L} \bar{W}_{\alpha}^{(kl)}(\bar{u}) (\nabla a \bar{F}_{\alpha}^{(k)}, \Delta a_{\alpha}) (\nabla a \bar{F}_{\alpha}^{(l)}, \Delta a_{\alpha})$$

We now re-write it using a random variable of mean 0 (the e_α vector) and it covaraince as:

$$ar{J}[ar{u}] = \sum_{lpha=1}^{N} (e_lpha, V[e_lpha]^- e_lpha)$$

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Strong hypothesis 2

- The rank of the covariance of e_α is the same as the hypothesis and each e_α is a random variable (Gaussian and independent)
- So $\bar{J}[\bar{u}]$ is a χ^2 variable, so we apply the ol' rejection method
- The hypothesis can be rejected with significance level a% if: $\bar{J}[\bar{u}]>\chi^2_{\rm rN,a}$
- Since $\overline{J}[u]$ requires the true data, it is approximated with the residual given the measured data J[u]

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Weak hypothesis

- To a first approximation is: $\bar{J}[\hat{u}] = \bar{J}[\bar{u}] - (\Delta u, \bar{V}[\hat{u}]^{-}\Delta u)$
- The first part is a χ^2 variable.... and also the second part
- The expectation and variance are LOWER than the ones for the strong hypothesis:
 E[J[ū]] = rN, V[J[ū]] = 2rN
 E[J[û]] = rN n', V[J[û]] = 2(rN n')

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Weak hypothesis 2

- The residual for the optimal estimate is (whp) smaller than for \bar{u}
- This is because its obtained minimizing the residual
- This analysis can be used to test that the constraints F^(k) are satiesfied by some value u
- The hypothesis is rejected with significance value a% if $\bar{J}[\hat{u}]>\chi^2_{\rm rN-n',a}$
- The same approximation is made for the residual as before

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Noise level

- If a covariance matrix y multiplied by a constant c...
- \blacksquare ... the pseudo inverse is multiplied by 1/c
- This does not affect the value that minimizes the residual (only the residual scale)
- The covariance is expressed as: $V[a_{\alpha}] = \epsilon^2 V_0[a_{\alpha}]$

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Noise level 2

- In practical problems, V₀ can be predicted, but not ϵ
- We can first estimate $\hat{u}...$
- and later estimate the noise level:

$$\hat{\epsilon}^2 = \frac{J_0[\hat{u}]}{rN - n'}$$

• The weak hypothesis can be re-writen as:

$$rac{\hat{\epsilon}^2}{\epsilon^2} > rac{\chi^2_{rN-n',a}}{rN-n'}$$

Matlab example

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The covariance

- K. assumes that you posses the covariance of the original data... but can you compute it?
- Example:
 - Given image correspondences, compute camera motion

- Then triangulate and obtain 3D points
- Fit some geometric object (line, plane, whatever)
- How does the error propagate thorough this?