## Exercises Measure Theoretic Probability

## Chapter 1

1. Prove the following statements.
(a) The intersection of an arbitrary family of $d$-systems is again a $d$ system.
(b) The intersection of an arbitrary family of $\sigma$-algebras is again a $\sigma$ algebra.
(c) If $\mathcal{C}_{1}$ and $\mathcal{C}_{2}$ are collections of subsets of $\Omega$ with $\mathcal{C}_{1} \subset \mathcal{C}_{2}$, then $d\left(\mathcal{C}_{1}\right) \subset$ $d\left(\mathcal{C}_{2}\right)$.
2. Let $\mathcal{G}$ and $\mathcal{H}$ be two $\sigma$-algebras on $\Omega$. Let $\mathcal{C}=\{G \cap H: G \in \mathcal{G}, H \in \mathcal{H}\}$. Show that $\mathcal{C}$ is a $\pi$-system and that $\sigma(\mathcal{C})=\sigma(\mathcal{G}, \mathcal{H})$.
3. Show that $\mathcal{D}_{2}$ (Williams, page 194) is a $\pi$-system.
4. Let $\Omega$ be a countable set. Let $\mathcal{F}=2^{\Omega}$ and let $p: \Omega \rightarrow[0,1]$ satisfy $\sum_{\omega \in \Omega} p(\omega)=1$. Put $\mathbb{P}(A)=\sum_{\omega \in A} p(\omega)$ for $A \in \mathcal{F}$. Show that $\mathbb{P}$ is a probability measure.
5. Let $\Omega$ be a countable set. Let $\mathcal{A}$ be the collection of $A \subset \Omega$ such that $A$ or its complement has finite cardinality. Show that $A$ is an algebra. What is $d(\mathcal{A})$ ?
6. Show that a finitely additive map $\mu: \Sigma_{0} \rightarrow[0, \infty]$ is countably additive if $\mu\left(H_{n}\right) \rightarrow 0$ for every decreasing sequence of sets $H_{n} \in \Sigma_{0}$ with $\bigcap_{n} H_{n}=\emptyset$. If $\mu$ is countably additive, do we necessarily have $\mu\left(H_{n}\right) \rightarrow 0$ for every decreasing sequence of sets $H_{n} \in \Sigma_{0}$ with $\bigcap_{n} H_{n}=\emptyset$ ?

## Chapter 2

1. Exercise of section 2.9 (page 28).

## Chapter 3

1. If $h_{1}$ and $h_{2}$ are measurable functions, then $h_{1} h_{2}$ is measurable too.
2. Let $X$ be a random variable. Show that $\Pi(X):=\left\{X^{-1}(-\infty, x]: x \in \mathbb{R}\right\}$ is a $\pi$-system and that it generates $\sigma(X)$.
3. Let $\left\{Y_{\gamma}: \gamma \in C\right\}$ be an arbitrary collection of random variables and $\left\{X_{n}: n \in \mathbb{N}\right\}$ be a countable collection of random variables, all defined on the same probability space.
(a) Show that $\sigma\left\{Y_{\gamma}: \gamma \in C\right\}=\sigma\left\{Y_{\gamma}^{-1}(B): \gamma \in C, B \in \mathcal{B}\right\}$.
(b) Let $\mathcal{X}_{n}=\sigma\left\{X_{1}, \ldots, X_{n}\right\}(n \in \mathbb{N})$ and $\mathcal{A}=\bigcup_{n=1}^{\infty} \mathcal{X}_{n}$. Show that $\mathcal{A}$ is an algebra and that $\sigma(\mathcal{A})=\sigma\left\{X_{n}: n \in \mathbb{N}\right\}$.
4. Show that the $X^{+}$and $X^{-}$are measurable functions and that $X^{+}$is right-continuous and $X^{-}$is left-continuous (notation as in section 3.12).
5. Let $\mathcal{F}$ be a $\sigma$-algebra on $\Omega$ with the property that for all $F \in \mathcal{F}$ it holds that $\mathbb{P}(F) \in\{0,1\}$. Let $X: \Omega \rightarrow \mathbb{R}$ be $\mathcal{F}$-measurable. Show that for some $c \in \mathbb{R}$ one has $\mathbb{P}(X=c)=1$. (Hint: $\mathbb{P}(X \leq x) \in\{0,1\}$ for all $x$.)

## Chapter 4

1. Williams, exercise E4.1.
2. Williams, exercise E4.6.
3. Let $\mathcal{G}_{1}, \mathcal{G}_{2}, \ldots$ be $\sigma$-algebras and let $\mathcal{G}=\sigma\left(\mathcal{G}_{1} \cup \mathcal{G}_{2} \cup \ldots\right)$.
(a) Show that $\Pi=\left\{G_{i_{1}} \cap G_{i_{2}} \cap \ldots \cap G_{i_{k}}: k \in \mathbb{N}, i_{k} \in \mathbb{N}, G_{i_{j}} \in \mathcal{G}_{i_{j}}\right\}$ is a $\pi$-system that generates $\mathcal{G}$.
(b) Assume that $\mathcal{G}_{1}, \mathcal{G}_{2}, \ldots$ is an independent sequence. Let $M$ and $N$ be disjoint subsets of $\mathbb{N}$ and put $\mathcal{M}=\sigma\left(\mathcal{G}_{i}, i \in M\right)$ and $\mathcal{N}=\sigma\left(\mathcal{G}_{i}, i \in\right.$ $N)$. Show that $\mathcal{M}$ and $\mathcal{N}$ are independent $\sigma$-algebras.

## Chapter 5

1. Show that the integral is a linear operator on $\mathcal{L}^{1}(S, \Sigma, \mu)$ by showing first that the result of section 5.5 holds true and then the general case.
2. Prove the second part of Scheffé's lemma (see page 55).
3. Consider a measure space $(S, \Sigma, \mu)$. Let $f \in(m \Sigma)^{+}$and define $\nu(E)=$ $\int_{S} f 1_{E} d \mu, E \in \Sigma$. Show that $\nu$ is a measure on $\Sigma$. Show also that $h \in \mathcal{L}^{1}(S, \Sigma, \nu)$ iff $f h \in \mathcal{L}^{1}(S, \Sigma, \mu)$ and that $\int_{S} h d \nu=\int_{S} f h d \mu$. (Use the 'standard machine' of section 5.12).
4. Let $\left(x_{1}, x_{2}, \ldots\right)$ be a sequence of nonnegative real numbers, let $\ell: \mathbb{N} \rightarrow \mathbb{N}$ be a bijection and define the sequence $\left(y_{1}, y_{2}, \ldots\right)$ by $y_{k}=x_{\ell(k)}$. Let for each $n$ the $n$-vector $y^{n}$ be given by $y^{n}=\left(y_{1}, \ldots, y_{n}\right)$. Consider then for each $n$ a sequence of numbers $x^{n}$ defined by $x_{k}^{n}=x_{k}$ if $x_{k}$ is a coordinate of $y^{n}$. Otherwise put $x_{k}^{n}=0$. Show that $x_{k}^{n} \uparrow x_{k}$ for every $k$ as $n \rightarrow \infty$. Show that $\sum_{k=1}^{\infty} y_{k}=\sum_{k=1}^{\infty} x_{k}$.
5. In this exercise $\lambda$ denotes Lebesgue measure on the Borel sets of $[0,1]$. Let $f:[0,1] \rightarrow \mathbb{R}$ be continuous. Then the Riemann integral $I:=\int_{0}^{1} f(x) d x$ exists (this is standard Analysis). But also the Lebesgue integral of $f$ exists. (Explain why.). Also explain why (use the definition of the Riemann
integral) there is a decreasing sequence of simple functions $U_{n}$ with limit $U$ satisfying $U \geq f$ and $\lambda\left(U^{n}\right) \downarrow I$. Prove that $\lambda(f)=I$.

## Chapter 6

1. Let $X$ and $Y$ be simple random variables. Show that $\mathbb{E} X$ doesn't depend on the chosen representation of $X$. Show also that $\mathbb{E}(X+Y)=\mathbb{E} X+\mathbb{E} Y$.
2. Show that for $X \in \mathcal{L}^{1}(\Omega, \mathcal{F}, \mathbb{P})$ it holds that $|\mathbb{E} X| \leq \mathbb{E}|X|$.
3. Let $X \in \mathcal{L}^{1}(\Omega, \mathcal{F}, \mathbb{P})$. Show that $\lim _{n \rightarrow \infty} n \mathbb{P}(|X|>n)=0$.
4. Prove the assertions (a)-(c) of section 6.5.
5. Complete the proof of theorem 6.11: show the a.s. uniqueness of $Y$ and show that if $X-Y \perp Z$ for all $Z$ in $\mathcal{K}$, then $\|X-Y\|_{2}=\inf \left\{\left\|X-Y^{\prime}\right\|_{2}\right.$ : $\left.Y^{\prime} \in \mathcal{K}\right\}$.
6. Prove lemma 6.12 with the 'standard machine' of section 5.12. With notation as in lemma 6.12, let $Y=h(X)$. Show also the following equality: $\mathbb{E} Y=\int_{\mathbb{R}} y \Lambda_{Y}(d y)$, with $\Lambda_{Y}$ the law of $Y$.

## Chapter 8

1. Prove part (b) of Fubini's theorem in section 8.2 for $f \in \mathcal{L}^{1}(S, \Sigma, \mu)$ (you already know it for $f \in \mathrm{~m} \Sigma^{+}$). Explain why $s_{1} \mapsto f\left(s_{1}, s_{2}\right)$ is in $\mathcal{L}^{1}\left(S_{1}, \Sigma_{1}, \mu_{1}\right)$ for all $s_{2}$ outside a set $N$ of $\mu_{2}$-measure zero and that $I_{2}^{f}$ is well defined on $N^{c}$.
2. If $Z_{1}, Z_{2}, \ldots$ is a sequence of nonnegative random variables, then

$$
\begin{equation*}
\mathbb{E} \sum_{k=1}^{\infty} Z_{k}=\sum_{k=1}^{\infty} \mathbb{E} Z_{k} \tag{1}
\end{equation*}
$$

Show that this follows from Fubini's theorem (as an alternative to section 6.5). If $\sum_{k=1}^{\infty} \mathbb{E} Z_{k}<\infty$, what is $\mathbb{P}\left(\sum_{k=1}^{\infty} Z_{k}=\infty\right)$. Formulate a result similar to (1) for random variables $Z_{k}$ that may assume negative values as well.
3. Let the vector of random variables $(X, Y)$ have a joint probability density function $f$. Let $f_{X}$ and $f_{Y}$ be the (marginal) probability density functions of $X$ and $Y$ respectively. Show that $X$ and $Y$ are independent iff $f(x, y)=$ $f_{X}(x) f_{Y}(y)$ for all $x, y$ except in a set of Leb $\times$ Leb-measure zero.
4. Let $f$ be defined on $\mathbb{R}^{2}$ such that for all $a \in \mathbb{R}$ the function $y \mapsto f(a, y)$ is Borel and such that for all $b \in \mathbb{R}^{2}$ the function $x \mapsto f(x, b)$ is continuous.

Show that for all $a, b, c \in \mathbb{R}$ the function $(x, y) \mapsto b x+c f(a, y)$ is Borelmeasurable on $\mathbb{R}^{2}$. Let $a_{i}^{n}=i / n, i \in \mathbb{Z}, n \in \mathbb{N}$. Define

$$
f^{n}(x, y)=\sum_{i} 1_{\left(a_{i-1}^{n}, a_{i}^{n}\right]}(x)\left(\frac{a_{i}^{n}-x}{a_{i}^{n}-a_{i-1}^{n}} f\left(a_{i-1}^{n}, y\right)+\frac{x-a_{i-1}^{n}}{a_{i}^{n}-a_{i-1}^{n}} f\left(a_{i}^{n}, y\right)\right) .
$$

Show that the $f^{n}$ are Borel-measurable on $\mathbb{R}^{2}$ and conclude that $f$ is Borel-measurable on $\mathbb{R}^{2}$.
5. Show that for $t>0$

$$
\int_{0}^{\infty} \sin x e^{-t x} \mathrm{~d} x=\frac{1}{1+t^{2}}
$$

Show that $x \mapsto \frac{\sin x}{x}$ is not in $\mathcal{L}^{1}(\mathbb{R}, \mathcal{B}$, Leb $)$, but that we can use Fubini's theorem to prove that the Riemann integral

$$
\int_{0}^{\infty} \frac{\sin x}{x} \mathrm{~d} x=\frac{\pi}{2}
$$

6. Let $F, G: \mathbb{R} \rightarrow \mathbb{R}$ be nondecreasing and right-continuous. Use Fubini's theorem to show the integration by parts formula, valid for all $a<b$,

$$
F(b) G(b)-F(a) G(a)=\int_{(a, b]} F(s-) d G(s)+\int_{(a, b]} G(s) d F(s)
$$

Hint: integrate $1_{(a, b]^{2}}$ and split the square into a lower and an upper triangle.
7. Let $F$ be the distribution function of a nonnegative random variable $X$ and assume that $\mathbb{E} X^{\alpha}<\infty$ for some $\alpha>0$. Use exercise 6 to show that

$$
\mathbb{E} X^{\alpha}=\alpha \int_{0}^{\infty} x^{\alpha-1}(1-F(x)) d x
$$

## Chapter 9

1. Finish the proof of theorem 9.2: Take arbitrary $X \in \mathcal{L}^{1}(\Omega, \mathcal{F}, \mathbb{P})$ and show that the existence of the conditional expectation of $X$ follows from the existence of the conditional expectations of $X^{+}$and $X^{-}$.
2. Prove the conditional version of Fatou's lemma, property (f) on page 88 (Williams).
3. Prove the conditional Dominated Convergence theorem, property (g) on page 88 (Williams).
4. Let $(X, Y)$ have a bivariate normal distribution with $\mathbb{E} X=\mu_{X}, \mathbb{E} Y=$ $\mu_{Y}, \operatorname{Var} X=\sigma_{X}^{2}, \operatorname{Var} Y=\sigma_{Y}^{2}$ and $\operatorname{Cov}(X, Y)=c$. Let

$$
\hat{X}=\mu_{x}+\frac{c}{\sigma_{Y}^{2}}\left(Y-\mu_{Y}\right)
$$

Show that $\mathbb{E}(X-\hat{X}) Y=0$. Show also (use a special property of the bivariate normal distribution) that $\mathbb{E}(X-\hat{X}) g(Y)=0$ if $g$ is a Borelmeasurable function such that $\mathbb{E} g(Y)^{2}<\infty$. Conclude that $\hat{X}$ is a version of $\mathbb{E}[X \mid Y]$.
5. Williams, Exercise E9.1.
6. Williams, Exercise E9.2
7. Let $X \in \mathcal{L}^{2}(\Omega, \mathcal{F}, \mathbb{P})$ and $Y \in \mathcal{L}^{2}(\Omega, \mathcal{G}, \mathbb{P})$, where $\mathcal{G}$ is a sub- $\sigma$-algebra of $\mathcal{F}$. Let $\hat{X}$ be a version of the conditional expectation $\mathbb{E}[X \mid \mathcal{G}]$. Show that

$$
\mathbb{E}(X-Y)^{2}=\mathbb{E}(X-\hat{X})^{2}+\mathbb{E}(Y-\hat{X})^{2}
$$

Deduce that $\hat{X}$ can be viewed as an orthogonal projection of $X$ onto $(\Omega, \mathcal{G}, \mathbb{P})$.

## Chapter 10

1. Let $X$ be an adapted process and $T$ a stopping time that is finite. Show that $X_{T}$ is $\mathcal{F}$-measurable. Show also that for arbitrary stopping times $T$ (so the value infinity is also allowed) the stopped process $X^{T}$ is adapted.
2. For every $n$ we have a measurable function $f_{n}$ on $\mathbb{R}^{n}$. Let $Z_{1}, Z_{2}, \ldots$ be independent random variables and $\mathcal{F}_{n}=\sigma\left(Z_{1}, \ldots, Z_{n}\right)$. Show that (you may assume sufficient integrability) that $X_{n}=f_{n}\left(Z_{1}, \ldots, Z_{n}\right)$ defines a martingale under the condition that $\mathbb{E} f_{n}\left(z_{1}, \ldots, z_{n-1}, Z_{n}\right)=f_{n-1}\left(z_{1}, \ldots, z_{n-1}\right)$ for every $n$.
3. If $S$ and $T$ are stopping times, then also $S+T, S \vee T$ and $S \wedge T$ are stopping times. Show this.
4. Show that an adapted process $X$ is a martingale iff $\mathbb{E}\left[X_{n+m} \mid \mathcal{F}_{n}\right]=X_{n}$ for all $n, m \geq 0$.
5. (a) If $X$ is a martingale is and $f$ a convex function such that $\mathbb{E}\left|f\left(X_{n}\right)\right|<$ $\infty$, then $Y$ defined by $Y_{n}=f\left(X_{n}\right)$ is a submartingale. Show this.
(b) Show that $Y$ is a submartingale, if $X$ is a submartingale and $f$ is a convex increasing function.
6. Prove Corollaries (c) and (d) on page 101.
7. Let $X_{1}, X_{2}, \ldots$ be an iid sequence of Bernoulli random variables. Put $\mathcal{F}_{n}=\sigma\left(X_{1}, \ldots, X_{n}\right), n \geq 1$. Let $M$ be a martingale adapted to the generated filtration. Show that the Martingale Representation Property holds: there exists a constant $m$ and a predictable process $Y$ such that $M_{n}=m+(Y \bullet X)_{n}, n \geq 1$.

## Chapter 11

1. Let $X$ be an adapted process and $a<b$ real numbers. Let $S_{1}=\inf \{n$ : $\left.X_{n}<a\right\}, T_{1}=\inf \left\{n>S_{1}: X_{n}>b\right\}$, etc. Show that the $S_{k}$ and $T_{k}$ are stopping times. Show also that the process $C$ of section 11.1 is previsible (synonymous for predictable).
2. Consider the probability space $(\Omega, \mathcal{F}, \mathbb{P})$ with $\Omega=[0,1), \mathcal{F}$ the Borel sets of $[0,1)$ and $\mathbb{P}$ the Lebesgue measure. Let $I_{k}^{n}=\left[k 2^{-n},(k+1) 2^{-n}\right)$ for $k=$ $0, \ldots, 2^{n}-1$ and $\mathcal{F}_{n}$ be the $\sigma$-algebra by the $I_{k}^{n}$ for $k=0, \ldots, 2^{n}-1$. Define $X_{n}=1_{I_{0}^{n}} 2^{n}$. Show that $X_{n}$ is a martingale is and that the conditions of theorem 11.5 are satisfied. What is $X_{\infty}$ in this case? Do we have $X_{n} \xrightarrow{\mathcal{L}^{1}} X_{\infty}$ ? (This has something to do with 11.6).
3. Let $X$ be a submartingale with $\sup _{n \geq 0} \mathbb{E}\left|X_{n}\right|<\infty$. Show that there exists a random variable $X_{\infty}$ such that $X_{n} \rightarrow X_{\infty}$ a.s.
4. Show that for a supermartingale $X$ the condition $\sup \left\{\mathbb{E}\left|X_{n}\right|: n \in \mathbb{N}\right\}<\infty$ is equivalent to the condition $\sup \left\{\mathbb{E} X_{n}^{-}: n \in \mathbb{N}\right\}<\infty$.

## Chapter 12

1. Exercise 12.1
2. Exercise 12.2
3. Let $\left(H_{n}\right)$ be a predictable sequence of random variables with $\mathbb{E} H_{n}^{2}<$ $\infty$ for all $n$. Let $\left(\varepsilon_{n}\right)$ be a sequence with $\mathbb{E} \varepsilon_{n}^{2}=1, \mathbb{E} \varepsilon_{n}=0$ and $\varepsilon_{n}$ independent of $\mathcal{F}_{n-1}$ for all $n$. Let $M_{n}=\sum_{k \leq n} H_{k} \varepsilon_{k}, n \geq 0$. Compute the conditional variance process $A$ of $\left(M_{n}\right)$. Take $p>1 / 2$ and consider $N_{n}=\sum_{k \leq n} \frac{1}{\left(1+A_{k}\right)^{p}} H_{k} \varepsilon_{k}$. Show that there exists a random variable $N_{\infty}$ such that $N_{n} \rightarrow N_{\infty}$ a.s. Show (use Kronecker's lemma) that $\frac{M_{n}}{\left(1+A_{n}\right)^{p}}$ has an a.s. finite limit.

## Chapter 13

1. Let $\mathcal{C}_{1}, \ldots, \mathcal{C}_{n}$ be uniformly integrable collections of random variables on a common probability space. Show that $\bigcup_{k=1}^{n} \mathcal{C}_{k}$ is uniformly integrable. (In particular is a finite collection in $\mathcal{L}^{1}$ uniformly integrable).
2. Williams, exercise E13.1.
3. Williams, exercise E13.2.
4. Let $\mathcal{C}$ be a uniformly integrable collection of random variables.
(a) Consider $\overline{\mathcal{C}}$, the closure of $\mathcal{C}$ in $\mathcal{L}^{1}$. Use E13.1 to show that also $\overline{\mathcal{C}}$ is uniformly integrable.
(b) Let $\mathcal{D}$ be the convex hull of $\mathcal{C}$, the smallest convex set that contains $\mathcal{C}$. Then both $\mathcal{D}$ and its closure in $\mathcal{L}^{1}$ are uniformly integrable
5. In this exercise you prove (fill in the details) the following characterization: a collection $\mathcal{C}$ is uniformly integrable iff there exists a function $G: \mathbb{R}^{+} \rightarrow$ $\mathbb{R}^{+}$such that $\lim _{t \rightarrow \infty} \frac{G(t)}{t}=\infty$ and $M:=\sup \{\mathbb{E} G(|X|): X \in \mathcal{C}\}<\infty$. The necessity you prove as follows. Let $\varepsilon>0$ choose $a=M / \varepsilon$ and $c$ such that $\frac{G(t)}{t} \geq a$ for all $t>c$. To prove uniform integrability of $\mathcal{C}$ you use that $|X| \leq \frac{G(|X|)}{a}$ on the set $\{|X| \geq c\}$.
It is less easy to prove sufficiency. Proceed as follows. Suppose that we have a sequence $\left(g_{n}\right)$ with $g_{0}=0$ and $\lim _{n \rightarrow \infty} g_{n}=\infty$. Define $g(t)=$ $\sum_{n=0}^{\infty} 1_{[n, n+1)}(t) g_{n}$ and $G(t)=\int_{0}^{t} g(s) d s$. Check that $\lim _{t \rightarrow \infty} \frac{G(t)}{t}=\infty$. With $a_{n}(X)=\mathbb{P}(|X|>n)$, it holds that $\mathbb{E} G(|X|) \leq \sum_{n=1}^{\infty} g_{n} a_{n}(|X|)$. Furthermore, for every $k \in \mathbb{N}$ we have $\int_{|X| \geq k}|X| d \mathbb{P} \geq \sum_{m=k}^{\infty} a_{m}(X)$. Pick for every $n$ a constant $c_{n} \in \mathbb{N}$ such that $\int_{|X| \geq c_{n}}|X| d \mathbb{P} \leq 2^{-n}$. Then $\sum_{m=c_{n}}^{\infty} a_{m}(X) \leq 2^{-n}$ and hence $\sum_{n=1}^{\infty} \sum_{m=c_{n}}^{\infty} a_{m}(X) \leq 1$. Choose then the sequence $\left(g_{n}\right)$ as the 'inverse' of $\left(c_{n}\right): g_{n}=\#\left\{k: c_{k} \leq n\right\}$.
6. Prove that a collection $\mathcal{C}$ is uniformly integrable iff there exists an increasing and convex function $G: \mathbb{R}^{+} \rightarrow \mathbb{R}^{+}$such that $\lim _{t \rightarrow \infty} \frac{G(t)}{t}=\infty$ and $M:=\sup \{\mathbb{E} G(|X|): X \in \mathcal{C}\}<\infty$. (You may use the result of exercise 5.) Let $\mathcal{D}$ be the closure of the convex hull of a uniformly integrable collection $\mathcal{C}$ in $\mathcal{L}^{1}$. With the function $G$ as above we have $\sup \{\mathbb{E} G(|X|): X \in \mathcal{D}\}=$ $M$, whence also $\mathcal{D}$ is uniformly integrable.
7. Let $p \geq 1$ and let $X, X_{1}, X_{2}, \ldots$ be random variables. Then $X_{n}$ converges to $X$ in $\mathcal{L}^{p}$ iff the following two conditions are satisfied.
(a) $X_{n} \rightarrow X$ in probability,
(b) The collection $\left\{\left|X_{n}\right|^{p}: n \in \mathbb{N}\right\}$ is uniformly integrable.
8. Exercise E13.3.

## Chapter 14

1. Let $Y \in \mathcal{L}^{1},\left(\mathcal{F}_{n}\right)$ and define for all $n \in \mathbb{N}$ the random variable $X_{n}=$ $\mathbb{E}\left[Y \mid \mathcal{F}_{n}\right]$. We know that there is $X_{\infty}$ such that $X_{n} \rightarrow X_{\infty}$ a.s. Show that for $Y \in \mathcal{L}^{2}$, we have $X_{n} \xrightarrow{\mathcal{L}^{2}} X_{\infty}$. Find a condition such that $X_{\infty}=Y$. Give also an example in which $P\left(X_{\infty}=Y\right)=0$.
2. Let $X=\left(X_{n}\right)_{n \leq 0}$ a (backward) supermartingale.
(a) Show equivalence of the next two properties:
(i) $\sup _{n} \mathbb{E}\left|X_{n}\right|<\infty$ and (ii) $\lim _{n \rightarrow-\infty} \mathbb{E} X_{n}<\infty$.
(Use that $x \mapsto x^{+}$is convex and increasing.)
(b) Under the condition $\sup _{n} \mathbb{E}\left|X_{n}\right|=: A<\infty$ the supermartingale $X$ is uniformly integrable. To show this, you may proceed as follows (but other solutions are equally welcome). Let $\varepsilon>0$ and choose $K \in \mathbb{Z}$ such that for all $n<K$ one has $0 \leq \mathbb{E} X_{n}-\mathbb{E} X_{K}<\varepsilon$. It is then sufficient to show that $\left(X_{n}\right)_{n \leq K}$ is uniformly integrable. Let $c>0$ be arbitrary and $F_{n}=\left\{\left|X_{n}\right|>c\right\}$. Using the supermartingale inequality you show that

$$
\int_{F_{n}}\left|X_{n}\right| d \mathbb{P} \leq \int_{F_{n}}\left|X_{K}\right| d \mathbb{P}+\varepsilon
$$

Because $\mathbb{P}\left(F_{n}\right) \leq \frac{A}{c}$ you conclude the proof.
3. Show that $R$, defined on page 142 of Williams, is equal to $q \mathbb{E}\left\|X^{p-1} Y\right\|$. Show also that the hypothesis of lemma 14.10 is true for $X \wedge n$ if it is true for $X$ and complete the proof of this lemma.
4. Exercise E14.1.
5. Exercise E14.2. Find the error in the statement of what you have to prove in (b).
6. Suppose that $\mathbb{Q}$ is a probability measure on $(\Omega, \mathcal{F})$ such that $\mathbb{Q} \ll \mathbb{P}$ with $d \mathbb{Q} / d \mathbb{P}=M_{\infty}$. Denote by $\mathbb{P}_{n}$ and $\mathbb{Q}_{n}$ the restrictions of $\mathbb{P}$ and $\mathbb{Q}$ to $\mathcal{F}_{n}$ $(n \geq 1)$. Show that $\mathbb{Q}_{n} \ll \mathbb{P}_{n}$ and that

$$
\frac{d \mathbb{Q}_{n}}{d \mathbb{P}_{n}}=M_{n}
$$

where $M_{n}=\mathbb{E}_{P}\left[M_{\infty} \mid \mathcal{F}_{n}\right]$.
7. Let $M$ be a nonnegative martingale with $\mathbb{E} M_{n}=1$ for all $n$. Define $\mathbb{Q}_{n}(F)=\mathbb{E} 1_{F} M_{n}$ for $F \in \mathcal{F}_{n}(n \geq 1)$. Show that for all $n$ and $k$ one has $\mathbb{Q}_{n+k}(F)=\mathbb{Q}_{n}(F)$ for $F \in \mathcal{F}_{n}$. Assume that $M$ is uniformly integrable. Show that there exists a probability measure $\mathbb{Q}$ on $\mathcal{F}_{\infty}=\sigma\left(\bigcup_{n} \mathcal{F}_{n}\right)$ that is absolutely continuous w.r.t. $\mathbb{P}$ and that is such that for all $n$ the restriction of $\mathbb{Q}$ to $\mathcal{F}_{n}$ coincides with $\mathbb{Q}_{n}$. Characterize $d \mathbb{Q} / d \mathbb{P}$.
8. Consider the set up of section 14.17 (Williams). Assume that

$$
\prod_{k=1}^{n} \mathbb{E}_{\mathbb{P}} \sqrt{\frac{g_{k}\left(X_{k}\right)}{f_{k}\left(X_{k}\right)}} \rightarrow 0
$$

Suppose one observes $X_{1}, \ldots, X_{n}$. Consider the testing problem $H_{0}$ : the densities of the $X_{k}$ are the $f_{k}$ against $H_{1}$ : the densities of the $X_{k}$ are the
$g_{k}$ and the test that rejects $H_{0}$ if $M_{n}>c_{n}$, where $\mathbb{P}\left(M_{n}>c_{n}\right)=\alpha \in(0,1)$ (likelihood ratio test). Show that this test is consistent: $\mathbb{Q}\left(M_{n} \leq c_{n}\right) \rightarrow 0$. (Side remark: the content of the Neyman-Pearson lemma is that this test is most powerful among all test with significance level less than or equal to $\alpha$.)
9. Finish the proof of theorem 14.11: Show that $\left\|Z_{n}\right\|_{p}$ is increasing in $n$ and that $\left\|Z_{\infty}\right\|_{p}=\sup \left\{\left\|Z_{n}\right\|_{p}: n \geq 1\right\}$.

## Chapter 17

1. Let $\mu, \mu_{1}, \mu_{2}, \ldots$ be probability measures on $\mathbb{R}$. Show that $\mu_{n} \xrightarrow{w} \mu$ iff for all bounded Lipschitz continuous functions one has $\int f d \mu_{n} \rightarrow \int f d \mu$. (Hint: for one implication the proof of lemma 17.2 is instructive.)
2. Show the 'if part' of lemma 17.2 without referring to the Skorohod representation. First you take for given $\varepsilon>0$ a $K>0$ such that $F(K)-$ $F(-K)>1-\varepsilon$. Approximate a continuous $f$ on the interval $(-K, K]$ with a piecewise constant function and you compute the integrals of this approximating function and use the convergence of the $F_{n}(x)$ at continuity points $x$ of $F$ etc.
3. If the random variables $X, X_{1}, X_{2}, \ldots$ are defined on the same probability space and if $X_{n} \xrightarrow{P} X$, then $X_{n} \xrightarrow{w} X$. Prove this.
4. Suppose that $X_{n} \xrightarrow{w} X$ and that the collection $\left\{X_{n}, n \geq 1\right\}$ is uniformly integrable (you make a minor change in the definition of this notion if the $X_{n}$ are defined on different probability spaces). Use the Skorohod representation to show that $X_{n} \xrightarrow{w} X$ implies $\mathbb{E} X_{n} \rightarrow \mathbb{E} X$.
5. Show the following variation on Fatou's lemma: if $X_{n} \xrightarrow{w} X$, then $\mathbb{E}|X| \leq$ $\liminf _{n \rightarrow \infty} \mathbb{E}\left|X_{n}\right|$.
6. Show that the weak limit of a sequence of probability measures is unique.
7. Look at the proof of the Helly-Bray lemma. You show that $F$ is rightcontinuous (use that for every $\varepsilon>0$ and $x \in \mathbb{R}$ there is a $c \in C$ such that $c>x$ and $F(x)>H(c)-\varepsilon$, take $y \in(x, c))$ and that $F_{n_{k}}(x)$ (the $n_{k}$ were obtained by the Cantor-type diagonalization procedure) converges to $F(x)$ at all continuity points $x$ of $F$ (take $c_{1}<x<c_{2}, c_{i} \in C$ and use that the $F_{n_{k}}\left(c_{i}\right)$ converge).
8. Consider the $N\left(\mu_{n}, \sigma_{n}^{2}\right)$ distributions, where the $\mu_{n}$ are real numbers and the $\sigma_{n}^{2}$ nonnegative. Show that this family is tight iff the sequences $\left(\mu_{n}\right)$ and $\left(\sigma_{n}^{2}\right)$ are bounded. Under what condition do we have that the $N\left(\mu_{n}, \sigma_{n}^{2}\right)$ distributions converge to a (weak) limit? What is this limit?

## Central limit theorem

1. For each $n$ we have a sequence $\xi_{n 1}, \ldots, \xi_{n k_{n}}$ of independent random variables with $\mathbb{E} \xi_{n j}=0$ and $\sum_{j=1}^{k_{n}} \operatorname{Var} \xi_{n j}=1$. If $\sum_{j=1}^{k_{n}} \mathbb{E}\left|\xi_{n j}\right|^{2+\delta} \rightarrow 0$ as $n \rightarrow \infty$ for some $\delta>0$, then $\sum_{j=1}^{k_{n}} \xi_{n j} \xrightarrow{w} N(0,1)$. Show that this follows from the Lindeberg Central Limit Theorem.
2. The classical central limit theorem says that $\frac{1}{\sigma \sqrt{n}} \sum_{j=1}^{n}\left(X_{j}-\mu\right) \xrightarrow{w} N(0,1)$, if the $X_{j}$ are iid with $\mathbb{E} X_{j}=\mu$ and $0<\operatorname{Var} X_{j}=\sigma^{2}<\infty$. Show that this follows from the Lindeberg Central Limit Theorem.
3. Show that $X_{n} \xrightarrow{w} X$ iff $\mathbb{E} f\left(X_{n}\right) \rightarrow \mathbb{E} f(X)$ for all bounded uniformly continuous functions $f$.
4. Let $X$ and $Y$ be independent, assume that $Y$ has a $N(0,1)$ distribution. Let $\sigma>0$. Let $\phi$ be the characteristic function of $X: \phi(u)=\mathbb{E} \exp (\mathrm{i} u X)$.
(a) Show that $Z=X+\sigma Y$ has density $p(z)=\frac{1}{\sigma \sqrt{2 \pi}} \mathbb{E} \exp \left(-\frac{1}{2 \sigma^{2}}(z-X)^{2}\right)$.
(b) Show that $p(z)=\frac{1}{2 \pi \sigma} \int \phi(-y / \sigma) \exp \left(\mathrm{i} y z / \sigma-\frac{1}{2} y^{2}\right) d y$.
5. Let $X, X_{1}, X_{2}, \ldots$ be a sequence of random variables and $Y$ a $N(0,1)$ distributed random variable independent of that sequence. Let $\phi_{n}$ be the characteristic function of $X_{n}$ and $\phi$ that of $X$. Let $p_{n}$ be the density of $X_{n}+\sigma Y$ and $p$ the density of $X+\sigma Y$.
(a) If $\phi_{n} \rightarrow \phi$ pointwise, then $p_{n} \rightarrow p$ pointwise. Invoke the previous exercise and the dominated convergence theorem to show this.
(b) Let $f \in C_{b}(\mathbb{R})$ be bounded by $B$. Show that $\mid \mathbb{E} f\left(X_{n}+\sigma Y\right)-\mathbb{E} f(X+$ $\sigma Y) \mid \leq 2 B \int\left(p(z)-p_{n}(z)\right)^{+} d z$.
(c) Show that $\left|\mathbb{E} f\left(X_{n}+\sigma Y\right)-\mathbb{E} f(X+\sigma Y)\right| \rightarrow 0$ if $\phi_{n} \rightarrow \phi$ pointwise.
(d) Prove the following theorem: $X_{n} \xrightarrow{w} X$ iff $\phi_{n} \rightarrow \phi$ pointwise.
6. Let $X_{1}, X_{2}, \ldots, X_{n}$ be an iid sequence having a distribution function $F$, a continuous density (w.r.t. Lebesgue measure) $f$. Let $m$ be such that $F(m)=\frac{1}{2}$. Assume that $f(m)>0$ and that $n$ is odd, $n=2 k-1$, say $\left(k=\frac{1}{2}(n+1)\right)$.
(a) Show that $m$ is the unique solution of the equation $F(x)=\frac{1}{2}$. We call $m$ the median of the distribution of $X_{1}$.
(b) Let $X_{(1)}=\min \left\{X_{1}, \ldots, X_{n}\right\}, X_{(2)}=\min \left\{X_{1}, \ldots, X_{n}\right\} \backslash\left\{X_{(1)}\right\}$, etc. The resulting $X_{(1)}, X_{(2)}, \ldots, X_{(n)}$ is called the ordered sample. The sample median $M_{n}$ of $X_{1}, \ldots, X_{n}$ is by definition $X_{(k)}$. Show that with $U_{n j}=1_{\left\{X_{j} \leq m+n^{-1 / 2} x\right\}}$ we have

$$
\mathbb{P}\left(n^{1 / 2}\left(M_{n}-m\right) \leq x\right)=\mathbb{P}\left(\sum_{j} U_{n j} \geq k\right) .
$$

(c) Let $p_{n}=\mathbb{P} U_{n j}, b_{n}=\left(n p_{n}\left(1-p_{n}\right)\right)^{1 / 2}, \xi_{n j}=\left(U_{n j}-p_{n}\right) / b_{n}, Z_{n}=$ $\sum_{j=1}^{n} \xi_{n j}, t_{n}=\left(k-n p_{n}\right) / b_{n}$. Rewrite the probabilities in part 6 b as $\mathbb{P}\left(Z_{n} \geq t_{n}\right)$ and show that $t_{n} \rightarrow t:=-2 x f(m)$.
(d) Show that $\mathbb{P}\left(Z_{n} \geq t\right) \rightarrow 1-\Phi(t)$, where $\Phi$ is the standard normal distribution.
(e) Show that $\mathbb{P}\left(Z_{n} \geq t_{n}\right) \rightarrow \Phi(2 f(m) x)$ and conclude that the Central Limit Theorem for the sample median holds:

$$
2 f(m) n^{1 / 2}\left(M_{n}-m\right) \xrightarrow{w} N(0,1)
$$

## Brownian motion

1. Consider the sequence of 'tents' $\left(X^{n}\right)$, where $X_{t}^{n}=n t$ for $t \in\left[0, \frac{1}{2 n}\right]$, $X_{t}^{n}=1-n t$ for $t \in\left[\frac{1}{2 n}, \frac{1}{n}\right]$, and zero elsewhere (there is no randomness here). Show that all finite dimensional distributions of the $X^{n}$ converge, but $X^{n}$ does not converge in distribution.
2. Show that $\rho$ as in (1.1) defines a metric.
3. Suppose that the $\xi_{i}$ of section 4 of the lecture notes are iid normally distributed random variables. Use Doob's inequality to obtain $\mathbb{P}\left(\max _{j \leq n}\left|S_{j}\right|>\right.$ $\gamma) \leq 3 \gamma^{-4} n^{2}$.
4. Show that a finite dimensional projection on $C[0, \infty)$ (with the metric $\rho$ ) is continuous.
5. Consider $C[0, \infty)$ with the Borel $\sigma$-algebra $\mathcal{B}$ induced by $\rho$ and some probability space $(\Omega, \mathcal{F}, \mathbb{P})$. If $X:(\Omega, \mathcal{F}) \rightarrow(C[0, \infty), \mathcal{B})$ is measurable, then all maps $\omega \mapsto X_{t}(\omega)$ are random variables. Show this, as well as its converse. For the latter you need separability that allows you to say that the Borel $\sigma$-algebra $\mathbb{B}$ is a product $\sigma$-algebra (see also Williams, page 82 ).
6. Prove proposition 2.2 of the lecture notes.
